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Expected Norms of Zero-One Polynomials

Peter Borwein, Kwok-Kwong Stephen Choi, and Idris Mercer

Abstract. Let $A_n = \{a_0 + a_1z + \cdots + a_{n-1}z^{n-1} : a_j \in \{0,1\}\}$, whose elements are called zeroone polynomials and correspond naturally to the 2^n subsets of $[n] := \{0, 1, \dots, n-1\}$. We also let $A_{n,m} = \{\alpha(z) \in A_n : \alpha(1) = m\}$, whose elements correspond to the $\binom{n}{m}$ subsets of [n] of size m, and let $\mathcal{B}_n = \mathcal{A}_{n+1} \setminus \mathcal{A}_n$, whose elements are the zero-one polynomials of degree exactly n.

Many researchers have studied norms of polynomials with restricted coefficients. Using $\|\alpha\|_p$ to denote the usual L_p norm of α on the unit circle, one easily sees that $\alpha(z) = a_0 + a_1 z + \cdots + a_N z^N \in \mathbb{R}[z]$ satisfies $\|\alpha\|_2^2 = c_0$ and $\|\alpha\|_4^4 = c_0^2 + 2(c_1^2 + \dots + c_N^2)$, where $c_k := \sum_{j=0}^{N-k} a_j a_{j+k}$ for $0 \le k \le N$.

If $\alpha(z) \in \mathcal{A}_{n,m}$, say $\alpha(z) = z^{\beta_1} + \dots + z^{\beta_m}$ where $\beta_1 < \dots < \beta_m$, then c_k is the number of times kappears as a difference $\beta_i - \beta_i$. The condition that $\alpha \in \mathcal{A}_{n,m}$ satisfies $c_k \in \{0,1\}$ for $1 \le k \le n-1$ is thus equivalent to the condition that $\{\beta_1,\ldots,\beta_m\}$ is a *Sidon set* (meaning all differences of pairs of elements are distinct).

In this paper, we find the average of $\|\alpha\|_4^4$ over $\alpha \in \mathcal{A}_n$, $\alpha \in \mathcal{B}_n$, and $\alpha \in \mathcal{A}_{n,m}$. We further show that our expression for the average of $\|\alpha\|_4^4$ over $\mathcal{A}_{n,m}$ yields a new proof of the known result: if $m = o(n^{1/4})$ and B(n, m) denotes the number of Sidon sets of size m in [n], then almost all subsets of [n] of size m are Sidon, in the sense that $\lim_{n\to\infty} B(n,m)/\binom{n}{m}=1$.

Introduction and Statement of Main Result

We let A_n denote the set $\{a_0 + a_1z + \cdots + a_{n-1}z^{n-1} : a_j \in \{0, 1\} \text{ for all } j\}$, and we call the elements of A_n zero-one polynomials. There is a natural bijection between the 2^n polynomials in \mathcal{A}_n and the 2^n subsets of $[n] := \{0, 1, \dots, n-1\}$. Generally, if $\alpha(z) \in \mathcal{A}_n$, we define

 $m := \alpha(1) = \text{ the number of coefficients of } \alpha(z) \text{ that are } 1,$

and we write $\alpha(z) = z^{\beta_1} + z^{\beta_2} + \cdots + z^{\beta_m}$ where $\beta_1 < \beta_2 < \cdots < \beta_m$, so $\{\beta_1, \beta_1, \dots, \beta_m\}$ is the subset of [n] that corresponds to $\alpha(z)$. We let $\mathcal{A}_{n,m}$ denote the set $\{\alpha(z) \in \mathcal{A}_n : \alpha(1) = m\}$, so $|\mathcal{A}_{n,m}| = \binom{n}{m}$ and $\mathcal{A}_n = \mathcal{A}_{n,0} \cup \mathcal{A}_{n,1} \cup \cdots \cup \mathcal{A}_{n,n}$. We also define $\mathcal{B}_n := \mathcal{A}_{n+1} \setminus \mathcal{A}_n$, so \mathcal{B}_n consists of the 2^n zero-one polynomials of degree exactly *n*.

A recurring theme in the literature is the problem of finding a polynomial with "small" norm subject to some restriction on its coefficients. (See [3, Problem 26], [5, Problem 19], or [1, Ch. 4, 15].) In general, for

(1.1)
$$\alpha(z) = a_0 + a_1 z + \dots + a_N z^N \in \mathbb{R}[z],$$

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we define the usual L_p norms of $\alpha(z)$ on the unit circle:

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$$\|\alpha\|_p := \left(\frac{1}{2\pi} \int_0^{2\pi} \left|\alpha(e^{i\theta})\right|^p d\theta\right)^{1/p},$$

where $p \ge 1$ is real. The main result of this paper, which appears as Theorem 4.1 in Section 4, is that if $n \ge 4$ and $m \le n$, we have

$$\begin{split} \mathbf{E}_{\mathcal{A}_n}(\|\alpha\|_4^4) &= \frac{4n^3 + 42n^2 - 4n + 3 - 3(-1)^n}{96}, \\ \mathbf{E}_{\mathcal{A}_{n,m}}(\|\alpha\|_4^4) &= 2m^2 - m + \frac{2m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(2n^2 - 4n + 1 - (-1)^n)}{2n^{[4]}}, \\ \mathbf{E}_{\mathcal{B}_n}(\|\alpha\|_4^4) &= \frac{4n^3 + 66n^2 + 188n + 87 + 9(-1)^n}{96}, \end{split}$$

where $\mathbf{E}_{\Omega}(\|\alpha\|_4^4)$ denotes the average of $\|\alpha\|_4^4$ over the polynomials in Ω , and the notation $x^{[k]}$ is shorthand for $x(x-1)\cdots(x-k+1)$. This complements results of Newman and Byrnes [7], who found the average of $\|\alpha\|_4^4$ over the 2^n polynomials of the form

(1.2)
$$a_0 + a_1 z + \dots + a_{n-1} z^{n-1}, \quad a_j \in \{+1, -1\} \text{ for all } j,$$

and Borwein and Choi [2], who found (among other things) the average of $\|\alpha\|_6^6$ and $\|\alpha\|_8^8$ over the 2^n polynomials of the form (1.2), and the average of $\|\alpha\|_2^2$, $\|\alpha\|_4^4$, and $\|\alpha\|_6^6$ over the 3^n polynomials of the form

$$a_0 + a_1 z + \dots + a_{n-1} z^{n-1}$$
, $a_j \in \{+1, 0, -1\}$ for all j .

2 **Autocorrelation**

Notice that if α is of the form (1.1) and |z| = 1, we have

$$|\alpha(z)|^{2} = \alpha(z)\overline{\alpha(z)} = (a_{0} + a_{1}z + \dots + a_{N}z^{N}) \left(a_{0} + a_{1}\frac{1}{z} + \dots + a_{N}\frac{1}{z^{N}}\right)$$
$$= c_{N}\frac{1}{z^{N}} + \dots + c_{1}\frac{1}{z} + c_{0} + c_{1}z + \dots + c_{N}z^{N},$$

where the c_k are the so-called (*aperiodic*) *autocorrelations* of α , defined for $0 \le k \le N$ by $c_k := \sum_{j=0}^{N-k} a_j a_{j+k}$. Using the general fact that

$$\frac{1}{2\pi} \int_0^{2\pi} \left(b_{-M} \frac{1}{z^M} + \dots + b_{-1} \frac{1}{z} + b_0 + b_1 z + \dots + b_M z^M \right) d\theta = b_0, \quad (z = e^{i\theta}),$$

we see that for α of the form (1.1), we have

$$\|\alpha\|_2^2 = \frac{1}{2\pi} \int_0^{2\pi} \left(c_N \frac{1}{z^N} + \dots + c_1 \frac{1}{z} + c_0 + c_1 z + \dots + c_N z^N \right) d\theta = c_0, \quad (z = e^{i\theta}),$$

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and

$$\|\alpha\|_4^4 = \frac{1}{2\pi} \int_0^{2\pi} \left(c_N \frac{1}{z^N} + \dots + c_1 \frac{1}{z} + c_0 + c_1 z + \dots + c_N z^N \right)^2 d\theta$$

$$= c_N^2 + \dots + c_1^2 + c_0^2 + c_1^2 + \dots + c_N^2 = c_0^2 + 2(c_1^2 + \dots + c_N^2), \quad (z = e^{i\theta}).$$

We further observe that, say,

$$c_k^2 = \left(\sum_{j=0}^{N-k} a_j a_{j+k}\right)^2 = \sum_{i=0}^{N-k} a_i a_{i+k} \cdot \sum_{j=0}^{N-k} a_j a_{j+k} = \sum_{i=0}^{N-k} \sum_{j=0}^{N-k} a_i a_j a_{i+k} a_{j+k}$$
$$= \sum_{i=0}^{N-k} \sum_{j=0}^{N-k} f(i,j).$$

Noting that $f(i, j) := a_i a_j a_{i+k} a_{j+k}$ satisfies f(i, j) = f(j, i), we have

(2.2)
$$c_k^2 = \sum_{i=0}^{N-k} \sum_{j=0}^{N-k} f(i,j) = \sum_{i=0}^{N-k} f(i,i) + 2 \sum_{0 \le i < j \le N-k} f(i,j)$$
$$= \sum_{i=0}^{N-k} a_i^2 a_{i+k}^2 + 2 \sum_{0 \le i < j \le N-k} a_i a_j a_{i+k} a_{j+k}.$$

If $\alpha(z) = a_0 + \cdots + a_{n-1}z^{n-1} = z^{\beta_1} + \cdots + z^{\beta_m} \in \mathcal{A}_{n,m}$, then we have $c_0 = m$ and c_k is the number of j such that a_j and a_{j+k} are both 1 and is equal to the number of times k appears as a difference $\beta_i - \beta_j$. Thus $c_1 + \cdots + c_{n-1} = m(m-1)/2$, and since the c_k are nonnegative integers, we have

Author: Redundant (?) displayed lines deleted.

$$(2.3) c_1^2 + \dots + c_{n-1}^2 > c_1 + \dots + c_{n-1} = m(m-1)/2$$

with equality if and only if $c_k \in \{0, 1\}$ for $1 \le k \le m-1$, or in other words, if and only if all differences of pairs of elements of $\{\beta_1, \ldots, \beta_m\}$ are distinct. If all differences of pairs of elements of $\{\beta_1, \ldots, \beta_m\}$ are distinct, we call $\{\beta_1, \ldots, \beta_m\}$ a Sidon set.

Using (2.1), we see that (2.3) and $c_0 = m$ prove the following.

Proposition 2.1 For any $\alpha(z) = z^{\beta_1} + \cdots + z^{\beta_m} \in A_{n,m}$, we have $\|\alpha\|_4^4 \ge 2m^2 - m$, with equality if and only if $\{\beta_1, \ldots, \beta_m\}$ is a Sidon set.

We observe also that (2.3) implies that $c_1^2 + \cdots + c_{n-1}^2 - m(m-1)/2$ is a nonnegative integer, and is zero if and only if $\{\beta_1, \dots, \beta_m\}$ is Sidon.

3 Some Facts and Notation

If Ω denotes \mathcal{A}_n , \mathcal{B}_n , or $\mathcal{A}_{n,m}$, then we turn Ω into a probability space by giving each polynomial $\alpha \in \Omega$ equal weight $p(\alpha)$.

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Generally, we will denote a typical element of A_n or $A_{n,m}$ by

$$\alpha(z) = a_0 + a_1 z + \dots + a_{n-1} z^{n-1},$$

and denote a typical element of \mathcal{B}_n by $\alpha(z) = a_0 + a_1 z + \cdots + a_{n-1} z^{n-1} + z^n$. As in Section 1, if $\alpha \in A_{n,m}$, we also write

$$\alpha(z) = z^{\beta_1} + z^{\beta_2} + \dots + z^{\beta_m}$$

where $\beta_1 < \beta_2 < \cdots < \beta_m$.

If Ω is one of the three spaces A_n , B_n , or $A_{n,m}$ and X is a random variable on Ω , we of course have $\mathbf{E}_{\Omega}(X) = \sum_{\alpha \in \Omega} X(\alpha) p(\alpha)$, and we will sometimes omit the subscript Ω if it is clear from the context what probability space we are considering.

Two facts we will use that are each immediate from first principles are Markov's inequality, Pr[X > a] < E(X)/a, where X is a nonnegative real random variable, and linearity of expectation, $\mathbf{E}(X_1 + \cdots + X_k) = \mathbf{E}(X_1) + \cdots + \mathbf{E}(X_k)$, which holds regardless of dependence or independence of the X_i .

4 Calculation of $E(\|\alpha\|_{A}^{4})$

Let j_1, j_2, j_3, j_4 denote distinct integers. We begin this section by finding some averages of products of a_{j_i} that we will need later. First, suppose our probability space Ω is A_n . We then have

(4.1)
$$\mathbf{E}(a_{j_1}a_{j_2}) = \frac{1}{2^n} (\text{number of } \alpha \in \mathcal{A}_n \text{ such that } a_{j_1} = a_{j_2} = 1) = \frac{2^{n-2}}{2^n} = \frac{1}{4},$$

and by similar reasoning, we have

(4.2)
$$\mathbf{E}(a_{i_1}a_{i_2}a_{i_3}) = 1/8, \quad \mathbf{E}(a_{i_1}a_{i_2}a_{i_3}a_{i_4}) = 1/16.$$

Now suppose our probability space Ω is $A_{n,m}$. We then have

(4.3)
$$\mathbf{E}(a_{j_1}a_{j_2}) = \frac{1}{\binom{n}{m}} (\text{number of } \alpha \in \mathcal{A}_{n,m} \text{ such that } a_{j_1} = a_{j_2} = 1)$$
$$= \frac{\binom{n-2}{m-2}}{\binom{n}{m}} = \frac{m(m-1)}{n(n-1)} = \frac{m^{[2]}}{n^{[2]}},$$

and by similar reasoning, we have

(4.4)
$$\mathbf{E}(a_{j_1}a_{j_2}a_{j_3}) = m^{[3]}/n^{[3]}, \quad \mathbf{E}(a_{j_1}a_{j_2}a_{j_3}a_{j_4}) = m^{[4]}/n^{[4]}.$$

We note that we need $n \geq 4$ in order for all expressions in (4.3) and (4.4) to be defined. For $\Omega = A_{n,m}$, the case $n \leq 3$ will be treated separately.

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Now if Ω is either of the probability spaces A_n or $A_{n,m}$, then equation (2.2) gives

(4.5)
$$c_k^2 = \sum_{i=0}^{n-k-1} a_i a_{i+k} + 2 \sum_{0 \le i < j \le n-k-1} a_i a_{i+k} a_j a_{j+k}.$$

We define $\lambda := n - k$ and also define

(4.6)
$$S := \sum_{i=0}^{\lambda-1} a_i a_{i+k},$$

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(4.7)
$$T := \sum_{0 \le i < j \le \lambda - 1} a_i a_j a_{i+k} a_{j+k},$$

which of course implies $c_k^2 = S + 2T$. If k = 0, then $c_k^2 = m^2$. So if $\Omega = A_{n,m}$, we have simply $\mathbf{E}(c_0^2) = m^2$, whereas if $\Omega = \mathcal{A}_n$, we have

(4.8)
$$E(c_0^2) = \sum_{m=0}^n \frac{\binom{n}{m}}{2^n} m^2.$$

It is an exercise to see that the right side of (4.8) evaluates to $(n^2 + n)/4$. Alternatively, we may observe that c_0 has a binomial distribution with parameters n and 1/2, which implies

(4.9)
$$\mathbf{E}(c_0^2) = \mathbf{Var}(c_0) + \mathbf{E}(c_0)^2 = n \cdot \frac{1}{2} \cdot \frac{1}{2} + \left(n \cdot \frac{1}{2}\right)^2 = \frac{n^2 + n}{4}.$$

Having found $\mathbf{E}(c_0^2)$ for $\Omega = \mathcal{A}_{n,m}$ and for $\Omega = \mathcal{A}_n$, we now shift our attention to $\mathbf{E}(c_k^2)$ for $k \neq 0$.

Assume $k \neq 0$, and observe that (4.5), (4.6), and (4.7) (and linearity of expectation) give us

(4.10)
$$\mathbf{E}(c_k^2) = \mathbf{E}(S) + 2\mathbf{E}(T) = \sum_{i=0}^{\lambda-1} \mathbf{E}(a_i a_{i+k}) + 2 \sum_{0 \le i < j \le \lambda-1} \mathbf{E}(a_i a_j a_{i+k} a_{j+k}).$$

Since $k \neq 0$, each of the λ terms in the sum E(S) is of the form $E(a_{j_1}a_{j_2})$ where $j_1 \neq 0$ j_2 . We thus have

(4.11)
$$\mathbf{E}(S) = \begin{cases} \lambda/4 & \text{if } \Omega = \mathcal{A}_n, \\ \lambda m^{[2]}/n^{[2]} & \text{if } \Omega = \mathcal{A}_{n,m}, \end{cases}$$

by (4.1) and (4.2). As for the $\binom{\lambda}{2}$ terms in the sum E(T), each term is of the form $\mathbf{E}(a_i a_i a_{i+k} a_{j+k})$. Since $k \neq 0$ and i < j, the four subscripts i, j, i+k, j+k constitute either three distinct integers (if j = i + k) or four distinct integers (if $j \neq i + k$). If $\{i, j, i+k, j+k\}$ consists of three distinct integers j_1, j_2, j_3 where j_3 is the one that is

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"repeated", then, since $a_i \in \{0,1\}$ for all j, we have $\mathbf{E}(a_i a_j a_{i+k} a_{j+k}) = \mathbf{E}(a_{j_1} a_{j_2} a_{j_3}^2) =$ $\mathbf{E}(a_{j_1}a_{j_2}a_{j_3})$, whereas, of course, if $\{i, j, i+k, j+k\}$ consists of four distinct integers, then $\mathbf{E}(a_i a_j a_{i+k} a_{j+k})$ is of the form $\mathbf{E}(a_{j_1} a_{j_2} a_{j_3} a_{j_4})$. Therefore, we now ask the question: For which of the $\binom{\lambda}{2}$ terms in the sum $\mathbf{E}(T)$ does the set $\{i, j, i+k, j+k\}$ consist of only three distinct integers?

For some $i \in \{0, 1, ..., \lambda - 1\}$, there is exactly one j satisfying both $i < j \le \lambda - 1$ and j = i + k, and for other values of i, there is no such j. We will say that i is of "type I" if the former criterion holds, and is of "type II" if the latter criterion holds. An integer *i* is of type I if and only if $i + k < \lambda$, or equivalently, $i < \lambda - k = n - 2k$. If $n-2k \le 0$ (i.e., if $k \ge \lceil n/2 \rceil$), then i < n-2k never happens, i.e., no i is of type I and so all of the $\binom{\lambda}{2}$ terms in the sum $\mathbf{E}(T)$ are of the form $\mathbf{E}(a_{j_1}a_{j_2}a_{j_3}a_{j_4})$. On the other hand, if n - 2k > 0 (i.e., if $k < \lceil n/2 \rceil$), then $i < n - 2k = \lambda - k$ sometimes happens; namely, it happens if and only if i is one of the $\lambda - k$ integers $0, 1, \dots, \lambda - k - 1$. In that case, each of those $\lambda - k$ values of i is of type I, which implies that precisely $\lambda - k$ of the $\binom{\lambda}{2}$ terms in the sum $\mathbf{E}(T)$ are of the form $\mathbf{E}(a_{j_1}a_{j_2}a_{j_3})$ and the remaining terms are of the form $\mathbf{E}(a_{i_1}a_{i_2}a_{i_3}a_{i_4})$.

This implies that we have

$$\mathbf{E}(T) = \begin{cases} \binom{\lambda}{2} \mathbf{E}(a_{j_1} a_{j_2} a_{j_3} a_{j_4}) & \text{if } k \ge \lceil n/2 \rceil, \\ \binom{\lambda}{2} \mathbf{E}(a_{j_1} a_{j_2} a_{j_3} a_{j_4}) & \\ + (\lambda - k) \left[\mathbf{E}(a_{j_1} a_{j_2} a_{j_3}) - \mathbf{E}(a_{j_1} a_{j_2} a_{j_3} a_{j_4}) \right] & \text{if } k < \lceil n/2 \rceil. \end{cases}$$

Thus, if $\Omega = A_n$, then

$$\mathbf{E}(T) = \begin{cases} \binom{\lambda}{2} / 16 & \text{if } k \ge \lceil n/2 \rceil, \\ \binom{\lambda}{2} / 16 + (\lambda - k) / 16 & \text{if } k < \lceil n/2 \rceil, \end{cases}$$

and hence by (4.10) and (4.11),

$$\mathbf{E}(c_k^2) = \begin{cases} \lambda/4 + \lambda(\lambda - 1)/16 & \text{if } k \ge \lceil n/2 \rceil, \\ \lambda/4 + \lambda(\lambda - 1)/16 + 2(\lambda - k)/16 & \text{if } k < \lceil n/2 \rceil. \end{cases}$$

On the other hand, if $\Omega = A_{n,m}$, then

$$\mathbf{E}(T) = \begin{cases} \binom{\lambda}{2} m^{[4]} / n^{[4]} & \text{if } k \ge \lceil n/2 \rceil, \\ \binom{\lambda}{2} m^{[4]} / n^{[4]} + (\lambda - k) [m^{[3]} / n^{[3]} - m^{[4]} / n^{[4]}] & \text{if } k < \lceil n/2 \rceil, \end{cases}$$

and hence

$$\mathbf{E}(c_k^2) = \begin{cases} \lambda \frac{m^{[2]}}{n^{[2]}} + \lambda(\lambda - 1) \frac{m^{[4]}}{n^{[4]}} & \text{if } k \ge \lceil n/2 \rceil, \\ \lambda \frac{m^{[2]}}{n^{[2]}} + \lambda(\lambda - 1) \frac{m^{[4]}}{n^{[4]}} + 2(\lambda - k) \left[\frac{m^{[3]}}{n^{[3]}} - \frac{m^{[4]}}{n^{[4]}} \right] & \text{if } k < \lceil n/2 \rceil. \end{cases}$$

It then follows that if $\Omega = A_n$, we have

$$(4.12) \quad \mathbf{E}(c_1^2 + \dots + c_{n-1}^2) = \sum_{k=1}^{n-1} \left(\frac{\lambda}{4}\right) + \sum_{k=1}^{n-1} \left(\frac{\lambda(\lambda-1)}{16}\right) + \sum_{k=1}^{\lceil n/2 \rceil - 1} \left(\frac{2(\lambda-k)}{16}\right),$$

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whereas if $\Omega = A_{n,m}$, we have

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(4.13)
$$\mathbf{E}(c_1^2 + \dots + c_{n-1}^2) = \sum_{k=1}^{n-1} \left(\lambda \frac{m^{[2]}}{n^{[2]}}\right) + \sum_{k=1}^{n-1} \left(\lambda (\lambda - 1) \frac{m^{[4]}}{n^{[4]}}\right) + \sum_{k=1}^{\lceil n/2 \rceil - 1} \left(2(\lambda - k) \left[\frac{m^{[3]}}{n^{[3]}} - \frac{m^{[4]}}{n^{[4]}}\right]\right).$$

Recalling that λ is simply shorthand for n - k, it is straightforward to verify that

$$\sum_{k=1}^{n-1} \lambda = \frac{n(n-1)}{2}, \quad \sum_{k=1}^{n-1} (\lambda^2 - \lambda) = \frac{n(n-1)(n-2)}{3},$$

and that

$$\sum_{k=1}^{\lceil n/2 \rceil - 1} 2(\lambda - k) = \begin{cases} n(n-2)/2 & \text{if } n \text{ is even,} \\ (n-1)^2/2 & \text{if } n \text{ is odd.} \end{cases}$$

So, if $\Omega = A_n$, then from (4.12) we get

$$\mathbf{E}(c_1^2 + \dots + c_{n-1}^2) = \begin{cases} \frac{1}{4} \cdot \frac{n(n-1)}{2} + \frac{1}{16} \cdot \frac{n(n-1)(n-2)}{3} + \frac{1}{16} \cdot \frac{n(n-2)}{2} & \text{if } n \text{ is even,} \\ \frac{1}{4} \cdot \frac{n(n-1)}{2} + \frac{1}{16} \cdot \frac{n(n-1)(n-2)}{3} + \frac{1}{16} \cdot \frac{(n-1)^2}{2} & \text{if } n \text{ is odd,} \end{cases}$$
$$= \begin{cases} (2n^3 + 9n^2 - 14n)/96 & \text{if } n \text{ is even,} \\ (2n^3 + 9n^2 - 14n + 3)/96 & \text{if } n \text{ is odd,} \end{cases}$$

which, using (2.1) and (4.9), implies

$$\mathbf{E}(\|\alpha\|_4^4) = \begin{cases} \frac{n^2 + n}{4} + \frac{2n^3 + 9n^2 - 14n}{48} = \frac{2n^3 + 21n^2 - 2n}{48} & \text{if } n \text{ is even,} \\ \frac{n^2 + n}{4} + \frac{2n^3 + 9n^2 - 14n + 3}{48} = \frac{2n^3 + 21n^2 - 2n + 3}{48} & \text{if } n \text{ is odd,} \end{cases}$$

or equivalently

(4.14)
$$\mathbf{E}_{\mathcal{A}_n}(\|\alpha\|_4^4) = \frac{4n^3 + 42n^2 - 4n + 3 - 3(-1)^n}{96}.$$

On the other hand, if $\Omega = A_{n,m}$, then from (4.13) we get

$$\begin{split} \mathbf{E}(c_1^2 + \dots + c_{n-1}^2) \\ &= \begin{cases} \frac{m^{[2]}}{n^{[2]}} \cdot \frac{n(n-1)}{2} + \frac{m^{[4]}}{n^{[4]}} \cdot \frac{n(n-1)(n-2)}{3} + \left(\frac{m^{[3]}}{n^{[3]}} - \frac{m^{[4]}}{n^{[4]}}\right) \cdot \frac{n(n-2)}{2} & \text{if } n \text{ is even,} \\ \frac{m^{[2]}}{n^{[2]}} \cdot \frac{n(n-1)}{2} + \frac{m^{[4]}}{n^{[4]}} \cdot \frac{n(n-1)(n-2)}{3} + \left(\frac{m^{[3]}}{n^{[3]}} - \frac{m^{[4]}}{n^{[4]}}\right) \cdot \frac{(n-1)^2}{2} & \text{if } n \text{ is odd,} \end{cases} \\ &= \begin{cases} \binom{m}{2} + m^{[4]}/(3(n-3)) + m^{[3]}(n-m)(n^2-2n)/(2n^{[4]}) & \text{if } n \text{ is even,} \\ \binom{m}{2} + m^{[4]}/(3(n-3)) + m^{[3]}(n-m)(n^2-2n+1)/(2n^{[4]}) & \text{if } n \text{ is odd,} \end{cases} \end{split}$$

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which, using (2.1), implies

$$\mathbf{E}(\|\alpha\|_4^4) = \begin{cases} 2m^2 - m + \frac{2m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(n^2-2n)}{n^{[4]}} & \text{if } n \text{ is even,} \\ 2m^2 - m + \frac{2m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(n^2-2n+1)}{n^{[4]}} & \text{if } n \text{ is odd,} \end{cases}$$

or equivalently

$$(4.15) \ \mathbf{E}_{\mathcal{A}_{n,m}}(\|\alpha\|_4^4) = 2m^2 - m + \frac{2m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(2n^2 - 4n + 1 - (-1)^n)}{2n^{[4]}}.$$

Notice that if m is fixed and n approaches infinity, then $\mathbf{E}_{A_{n,m}}(\|\alpha\|_4^4)$ approaches $2m^2 - m$, i.e., for fixed m and large n, we expect a random $\alpha \in A_{n,m}$ to correspond to a Sidon set, as is consistent with intuition.

If
$$\Omega = \mathcal{B}_n$$
, since $\mathcal{B}_n := \mathcal{A}_{n+1} \setminus \mathcal{A}_n$, we get

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$$\begin{split} \mathbf{E}_{\mathcal{B}_n}(\|\alpha\|_4^4) &= \frac{1}{2^n} \sum_{\alpha \in \mathcal{B}_n} \|\alpha\|_4^4 = 2\mathbf{E}_{\mathcal{A}_{n+1}}(\|\alpha\|_4^4) - \mathbf{E}_{\mathcal{A}_n}(\|\alpha\|_4^4) \\ &= \frac{4n^3 + 66n^2 + 188n + 87 + 9(-1)^n}{96} \end{split}$$

by (4.14). Therefore we have proved

Theorem 4.1 If $m \le n$, we have

$$\begin{split} \mathbf{E}_{\mathcal{A}_n}(\|\alpha\|_4^4) &= \frac{4n^3 + 42n^2 - 4n + 3 - 3(-1)^n}{96}, \\ \mathbf{E}_{\mathcal{A}_{n,m}}(\|\alpha\|_4^4) &= 2m^2 - m + \frac{2m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(2n^2 - 4n + 1 - (-1)^n)}{2n^{[4]}} \\ &\qquad \qquad (if \ n \geq 4), \end{split}$$

$$\mathbf{E}_{\mathbb{B}_n}(\|\alpha\|_4^4) = \frac{4n^3 + 66n^2 + 188n + 87 + 9(-1)^n}{96}.$$

For completeness, we also determine $\mathbf{E}_{\mathcal{A}_{n,m}}(\|\alpha\|_4^4)$ when $n \leq 3$. If $n \leq 3$, we have $\alpha(z) = a_0 + a_1 z + a_2 z^2$ and then

$$\|\alpha\|_4^4 = c_0^2 + 2c_1^2 + 2c_2^2$$

$$= (a_0^2 + a_1^2 + a_2^2)^2 + 2(a_0a_1 + a_1a_2)^2 + 2(a_0a_2)^2$$

$$= a_0^4 + a_1^4 + a_2^4 + 4(a_0^2a_1^2 + a_0^2a_2^2 + a_1^2a_2^2) + 4a_0a_1^2a_2$$

$$= a_0 + a_1 + a_2 + 4(a_0a_1 + a_0a_2 + a_1a_2) + 4a_0a_1a_2,$$

since $a_i \in \{0, 1\}$ from which it readily follows that

$$\begin{split} \mathbf{E}_{\mathcal{A}_{2,0}}(\|\alpha\|_{4}^{4}) &= \mathbf{E}_{\mathcal{A}_{3,0}}(\|\alpha\|_{4}^{4}) = 0, \\ \mathbf{E}_{\mathcal{A}_{2,1}}(\|\alpha\|_{4}^{4}) &= \mathbf{E}_{\mathcal{A}_{3,1}}(\|\alpha\|_{4}^{4}) = 1, \\ \mathbf{E}_{\mathcal{A}_{2,2}}(\|\alpha\|_{4}^{4}) &= \mathbf{E}_{\mathcal{A}_{3,2}}(\|\alpha\|_{4}^{4}) = 6, \\ \mathbf{E}_{\mathcal{A}_{3,3}}(\|\alpha\|_{4}^{4}) &= 19. \end{split}$$

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We remark that substituting $m \in \{0, 1, 2, 3\}$ into the second equation in Theorem 4.1 and then formally cancelling common factors as appropriate, we get

$$\mathbf{E}_{\mathcal{A}_{n,3}}(\|\alpha\|_{4}^{4}) = 15 + \frac{3(2n^{2} - 4n + 1 - (-1)^{n})}{n(n-1)(n-2)},$$

$$\mathbf{E}_{\mathcal{A}_{n,2}}(\|\alpha\|_{4}^{4}) = 6,$$

$$\mathbf{E}_{\mathcal{A}_{n,1}}(\|\alpha\|_{4}^{4}) = 1,$$

$$\mathbf{E}_{A_{n,0}}(\|\alpha\|_{A}^{4})=0,$$

$$\mathbf{E}_{\mathcal{A}_{n,0}}(\|\alpha\|_4^4)=0$$

yielding results consistent with the explicit averages just obtained for $n \le 3$.

Ubiquity of Sidon Sets

We show that our expression for $\mathbf{E}_{A_{n,m}}(\|\alpha\|_4^4)$ yields a new proof of a result that appears in articles by Godbole et al. [4] and Nathanson [6].

Suppose $\Omega = A_{n,m}$, and as before, denote a typical element of $A_{n,m}$ by

$$\alpha(z)=z^{\beta_1}+\cdots+z^{\beta_m}.$$

Recall from Section 2 that $X := c_1^2 + \cdots + c_{n-1}^2 - {m \choose 2}$ is a nonnegative-integer-valued random variable on Ω that attains the value 0 if and only if $\{\beta_1, \ldots, \beta_m\}$ is a Sidon

We have

$$\mathbf{E}_{\mathcal{A}_{n,m}}(X) = \mathbf{E}_{\mathcal{A}_{n,m}}(c_1^2 + \dots + c_{n-1}^2) - \binom{m}{2}$$

$$= \begin{cases} \frac{m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(n^2-2n)}{2n^{[4]}} & \text{if } n \text{ is even,} \\ \frac{m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(n^2-2n+1)}{2n^{[4]}} & \text{if } n \text{ is odd} \end{cases}$$

$$\leq \frac{m^{[4]}}{3(n-3)} + \frac{m^{[3]}(n-m)(n-1)^2}{2n^{[4]}}$$

$$= \frac{m(m-1)(m-2)(2mn-3n-m)}{6n(n-2)}$$

$$\leq \frac{m^4}{3n}$$

if $n \ge 4$. On the other hand, if we let B(n, m) be the number of Sidon sets in [n] with *m* elements, then we have

$$\mathbf{E}(X) = \frac{1}{\binom{n}{m}} \sum_{\alpha \in \mathcal{A}_{n,m}} X = \frac{1}{\binom{n}{m}} \sum_{\alpha \in \mathcal{A}_{n,m}, X > 0} X \ge \frac{1}{\binom{n}{m}} \#\{\alpha \in \mathcal{A}_{n,m} : X(\alpha) > 0\}$$
$$\ge 1 - \frac{1}{\binom{n}{m}} B(n, m).$$

Hence we have proved (by essentially using Markov's inequality) the following.

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59___ 60___ **Corollary 5.1** For $4 \le m \le n$, we have

$$B(n,m) \ge \binom{n}{m} \left(1 - \frac{m^4}{3n}\right)$$

and

$$\Pr[\{\beta_1,\ldots,\beta_m\} \text{ is Sidon}] > 1 - \frac{m^4}{3n}.$$

Hence if $m = o(n^{1/4})$, then as n approaches infinity, the probability that a randomly chosen m-subset of [n] is Sidon approaches 1.

Although when $m = o(n^{1/4})$, the probability that a randomly chosen m-subset of [n] is Sidon approaches 1 (i.e., $\|\alpha\|_4^4$ is $2m^2 - m$ for almost all $\alpha \in A_{n,m}$), there are some other cases in which a positive proportion of polynomials in $A_{n,m}$ have very large L_4 norm.

For $\alpha \in A_{n,m}$, since for $0 \le k \le n-1$, $c_k = \sum_{j=0}^{n-k-1} a_j a_{j+k}$, we have $c_0 = m$, and for $1 \le k \le n-1$, $|c_k| \le \min\{m-1, n-k\}$. Therefore, we have

$$\|\alpha\|_4^4 = c_0^2 + 2\sum_{k=1}^{n-1} c_k^2 \le m^2 + 2\sum_{k=1}^{n-m+1} (m-1)^2 + 2\sum_{k=n-m+2}^{n-1} (n-k)^2$$

$$= 2nm^2 - \frac{4}{3}m^3 + 4m^2 - 4nm + 2n - \frac{5}{3}m$$

$$= 2(1+o(1))m^2 \left(n - \frac{2}{3}m\right)$$

if $n = o(m^2)$ as $m, n \to \infty$ and on the other hand, from (4.15) we have

$$\begin{split} \frac{2(1+o(1))m^4}{3n} &\leq \mathbf{E}_{\mathcal{A}_{n,m}}(\|\alpha\|_4^4) = \frac{1}{\binom{n}{m}} \sum_{\alpha \in \mathcal{A}_{n,m}} \|\alpha\|_4^4 \\ &= \frac{1}{\binom{n}{m}} \left\{ \sum_{\|\alpha\|_4^4 \leq x} \|\alpha\|_4^4 + \sum_{\|\alpha\|_4^4 > x} \|\alpha\|_4^4 \right\} \\ &\leq x + \frac{1}{\binom{n}{m}} \sum_{\|\alpha\|_4^4 > x} \|\alpha\|_4^4 \\ &\leq x + \frac{1}{\binom{n}{m}} \sum_{\|\alpha\|_4^4 > x} 2(1+o(1))m^2 \left(n - \frac{2}{3}m\right). \end{split}$$

It then follows that for any $x < 2(1 + o(1))m^4/(3n)$, we have

$$\frac{\#\{\alpha \in \mathcal{A}_{n,m} : \left\|\alpha\right\|_{4}^{4} > x\}}{\binom{n}{m}} \ge \frac{2(1 + o(1))m^{4}/(3n) - x}{2(1 + o(1))m^{2}(n - 2m/3)}.$$

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In particular, for any $\epsilon > 0$, if $m = c_1 n$ and $x = c_2 m^2 n$ for $0 < c_1 < 1$ and $0 < c_2 < 2(1 - \epsilon)c_1^2/3$, we have

$$\frac{\#\{\alpha \in \mathcal{A}_{n,m} : \|\alpha\|_4^4 > c_2 m^2 n\}}{\frac{n}{m}} \ge \frac{2(1-\epsilon)c_1^2/3 - c_2}{2(1+\epsilon)(1-2c_1/3)} > 0$$

for sufficiently large n and m, i.e., there is a positive proportion of polynomials in $A_{n,m}$ having large L_4 norm (note that the L_4 norm in $A_{n,m}$ is at most as large as $2(1+o(1))m^2n$.

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